Lecture 03: Classical Inferential Statistics I: Basics and Confidence Intervals

3.1. Ordinary Least Squares (OLS) Estimator: Expectation and Covariance

- Only stochasticity: residual errors $\boldsymbol{\epsilon}$ according to $\boldsymbol{y} = \mathbf{X}\,\boldsymbol{\beta} + \boldsymbol{\epsilon}$
- The OLS estimator is linear in y:

$$\hat{\beta} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'y$$

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&= \underline{\underline{\boldsymbol{\beta} + \left(\mathbf{X}'\mathbf{X}\right)^{-1}\mathbf{X}'\boldsymbol{\epsilon}}}
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Expectation value

$$E(\hat{\boldsymbol{\beta}}) = E(\boldsymbol{\beta}) + (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'E(\boldsymbol{\epsilon}) = \boldsymbol{\beta}$$

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The OLS estimator of parameter-linear models is **unbiased** under the mild condition $E(\boldsymbol{\epsilon}) = \mathbf{0}$ for all the data points

OLS estimator: variances and covariances

- Gauß-Markow conditions $\to \epsilon \sim$ i.i.d$N(0, \sigma^2) \to \hat{\beta}$ is normal distributed
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The variance-covariance matrix depends only on the values of the exogenous factors!

Results

- Ordinary least squares (OLS) estimator:

$$\boxed{\hat{\beta} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'y}$$

- **Variance-Covariance matrix** of the estimation errors (provided the errors are i.i.d.) can be written in terms of the **Hesse matrix H** of the objective function SSE:

$$\begin{aligned} \mathbf{V}_{\hat{\beta}} &= E\left((\hat{\beta}-\beta)(\hat{\beta}-\beta)'\right) = \sigma^2(\mathbf{X}'\mathbf{X})^{-1} = 2\sigma^2\mathbf{H}^{-1}, \\ H_{jk} &= \left.\frac{\partial^2 S}{\partial\beta_j\ \partial\beta_k}\right|_{\beta=\hat{\beta}} = 2(\mathbf{X}'\mathbf{X})_{jk} \end{aligned}$$

- Variances of estimation errors: $V(\hat{\beta}_j) = V_{jj}$
- Correlation of estimation errors: $\text{Corr}(\hat{\beta}_j, \hat{\beta}_k) = \frac{V_{jk}}{\sqrt{V_{jj}V_{kk}}}$
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Estimation of the residual variance

The above cannot be applied directly since the residual variance σ^2 is unknown and must be estimated by the minimum SSE $S(\hat{\beta})$:

$$\hat{\sigma}^2 = \frac{1}{n-J-1}\sum_i \left(y_i - \hat{y}(\boldsymbol{x}_i)\right)^2 = \frac{S(\hat{\beta})}{n-J-1}$$

Under the Gauß-Markow assumptions, this can be expressed as the sum of squared Gaussians as follows (derivation for the experts):

$$\begin{aligned}(n-J-1)\hat{\sigma}^2 &= (\hat{\boldsymbol{y}} - \boldsymbol{y})' (\hat{\boldsymbol{y}} - \boldsymbol{y}) \\ &= (\mathbf{X}\hat{\boldsymbol{\beta}} - \boldsymbol{y})'(\mathbf{X}\hat{\boldsymbol{\beta}} - \boldsymbol{y}) \\ &= (\mathbf{X}\hat{\boldsymbol{\beta}})'(\mathbf{X}\hat{\boldsymbol{\beta}}) - (\mathbf{X}\boldsymbol{\beta})'\boldsymbol{y} - \boldsymbol{y}'(\mathbf{X}\boldsymbol{\beta}) + \boldsymbol{y}'\boldsymbol{y}\end{aligned}$$

With following rule for scalar products: $\boldsymbol{a}'\boldsymbol{b} = \boldsymbol{b}'\boldsymbol{a}$ it follows that the two middle terms are equal. Replacing $\hat{\boldsymbol{\beta}} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\boldsymbol{y}$ we see that, interestingly, the first term is the negative of each of the two middle terms resulting in

$$(n-J-1)\hat{\sigma}^2 = \boldsymbol{y}' \left(\mathbf{1} - \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\right)\boldsymbol{y}$$

Estimation of the residual variance (ctned)

Finally, we replace the observed endogeneous data vector $\boldsymbol{y}$ by the model $\boldsymbol{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}$ Notice: the true and, according to the Gauß-Markow assumptions, immutable parameter vector $\boldsymbol{\beta}$ is used here!:

$$
\begin{aligned}
(n-J-1)\hat{\sigma}^2 &= (\mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon})' \left(\mathbf{1} - \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\right) (\mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}) \\
&= \boldsymbol{\epsilon}'(\mathbf{1} - \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}')\boldsymbol{\epsilon} \\
&\quad + 2(\mathbf{X}\boldsymbol{\beta})'(\mathbf{1} - \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}')\boldsymbol{\epsilon} + \boldsymbol{\beta}'\mathbf{X}'(\mathbf{1} - \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}
\end{aligned}
$$

After doing the simplification, we realize that the second and third term are each equal to zero, so we have the final result

$$
(n-J-1)\hat{\sigma}^2 = \boldsymbol{\epsilon}'(\mathbf{1} - \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}')\boldsymbol{\epsilon}
$$

With the Gauß-Markow-assumptions, this is proportional to a sum of $(n-J-1)$ squared Gaussians, i.e., a $\chi^2(n-J-1)$ distributed random variable

Results if the variance needs to be estimated

- Estimated variance-covariance matrix:

$$\hat{\mathbf{V}}_{\hat{\beta}} = 2\hat{\sigma}^2 \mathbf{H}^{-1} = \hat{\sigma}^2 \left(\mathbf{X}'\mathbf{X}\right)^{-1}$$

- The normalized approximate estimation errors are student-t distributed (a Gaussian in the numerator and the square root of a χ^2 distributed random variable in the denominator):

$$\frac{\hat{\beta}_j - \beta_j}{\sqrt{\hat{V}_{jj}}} \sim T(n-1-J)$$

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Multivariate distribution function of $\hat{\beta}$

The distribution of the errors $\Delta\hat{\boldsymbol{\beta}} = \hat{\boldsymbol{\beta}} - \boldsymbol{\beta}$ obeys a multivariate normal distribution:

$$f_{\hat{\boldsymbol{\beta}}}(\Delta\hat{\boldsymbol{\beta}}) \propto \exp\left[-\frac{1}{2}\Delta\hat{\boldsymbol{\beta}}' \,\mathbf{V}^{-1}\, \Delta\hat{\boldsymbol{\beta}}\right] = \exp\left[-\frac{\Delta\hat{\boldsymbol{\beta}}' \,\mathbf{X}'\mathbf{X}\, \Delta\hat{\boldsymbol{\beta}}}{2\sigma_\epsilon^2}\right].$$

Relation to the maximum-likelihood-method (→ Lecture 07:)

Expand the SSE $S(\beta)$ around $\hat{\beta}$ to second order:

$$S(\beta) - S(\hat{\beta}) \approx \frac{1}{2}\Delta\hat{\beta}' \,\mathbf{H}\, \Delta\hat{\beta} = \Delta\hat{\beta}' \,\mathbf{X}'\mathbf{X}\, \Delta\hat{\beta}$$

$$\Rightarrow \qquad \boxed{f_{\hat{\beta}}(\Delta\hat{\beta}) \propto \exp\left[-\frac{S(\beta) - S(\hat{\beta})}{2\sigma_\epsilon^2}\right]}$$

and with the estimated residual variance $\hat{\sigma}_\epsilon^2 = S(\hat{\beta})/(n-J-1)$

$$\hat{f}_{\hat{\beta}}(\beta) \propto \exp\left[-\frac{(n-J-1)}{2}\left(\frac{S(\beta)}{S(\hat{\beta})} - 1\right)\right]$$

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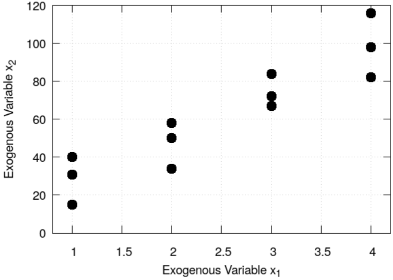
$$\Rightarrow \qquad \boxed{f_{\hat{\boldsymbol{\beta}}}(\Delta\hat{\boldsymbol{\beta}}) \propto \exp\left[-\frac{S(\boldsymbol{\beta}) - S(\hat{\boldsymbol{\beta}})}{2\sigma_\epsilon^2}\right]}$$

and with the estimated residual variance $\hat{\sigma}_\epsilon^2 = S(\hat{\boldsymbol{\beta}})/(n-J-1)$

$$\hat{f}_{\hat{\boldsymbol{\beta}}}(\boldsymbol{\beta}) \propto \exp\left[-\frac{(n-J-1)}{2}\left(\frac{S(\boldsymbol{\beta})}{S(\hat{\boldsymbol{\beta}})} - 1\right)\right]$$

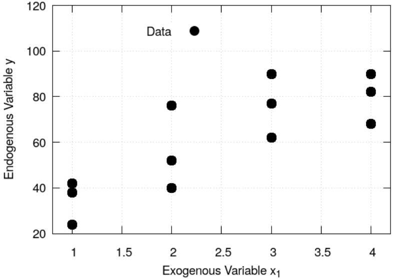
Example of correlated errors: modeling the demand for hotel rooms

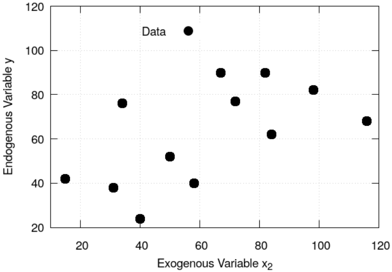


- The example of Lecture 02:

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \epsilon$$

- Exogenous factors: $x_0 = 1$, x_1: proxy for quality [# stars]; x_2: price [€/night].
- Endogenous variable: booking rate [%]
- The demand is positively correlated with both the quality and the price





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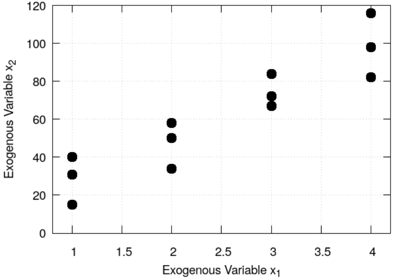

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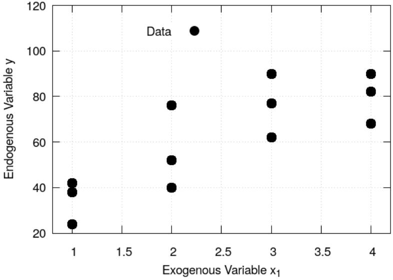

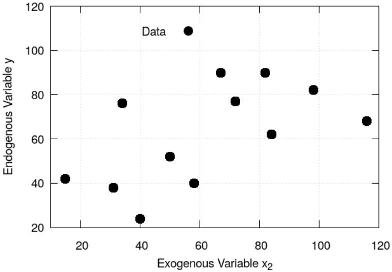

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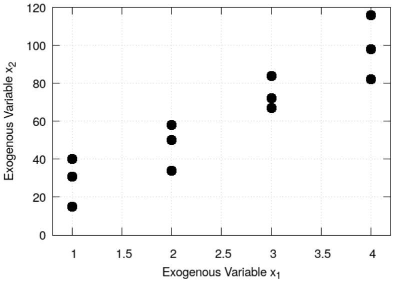

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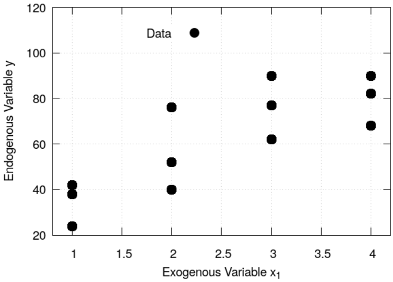

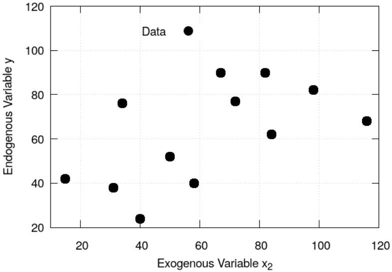

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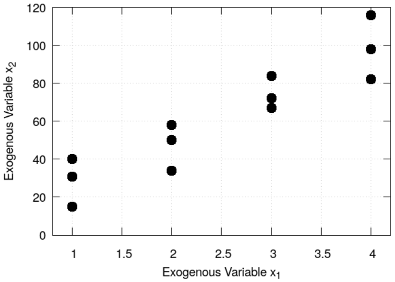

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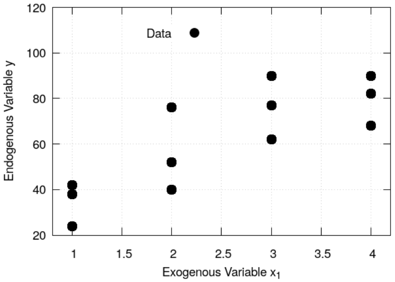
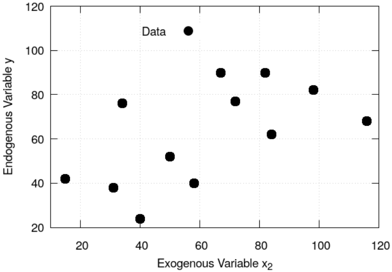
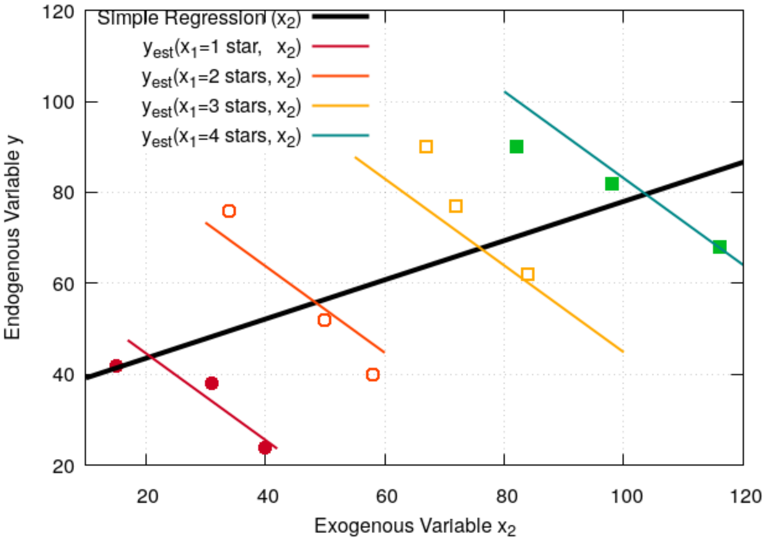
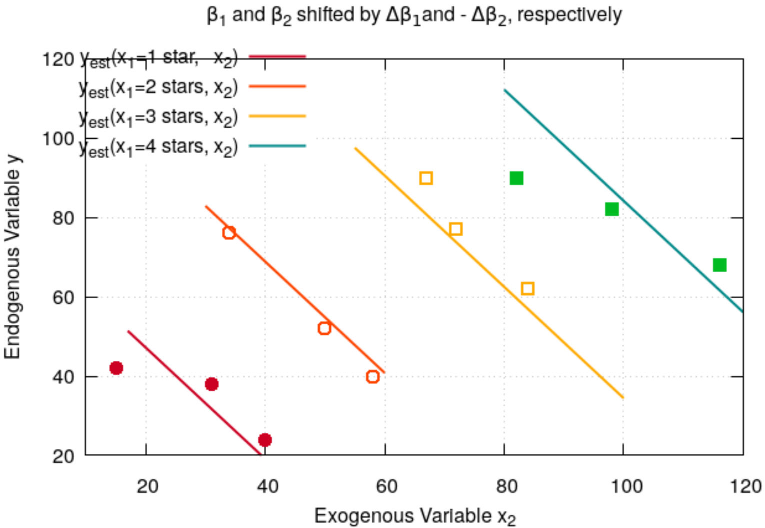




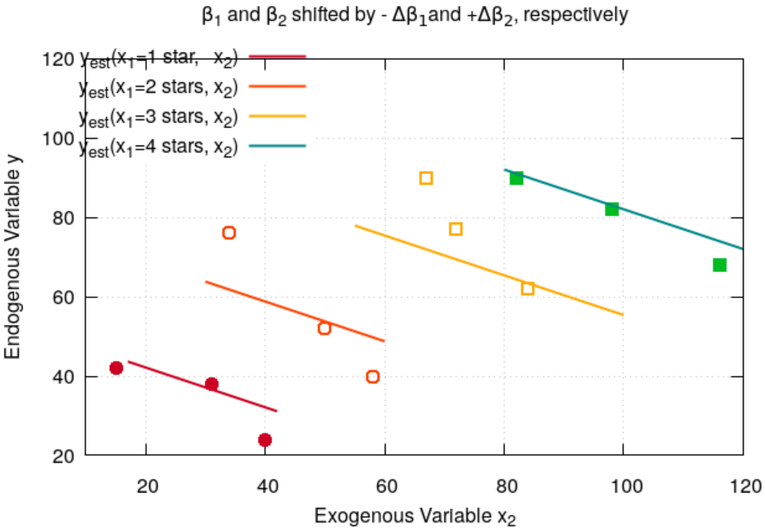
Residual errors for fitted parameters



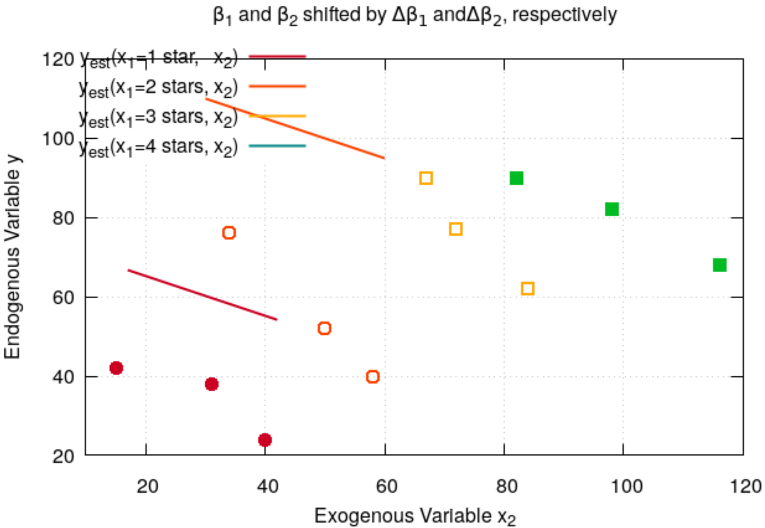
Effect of mis-fit parameters I: small effect if β_1 and β_2 have opposite misfits



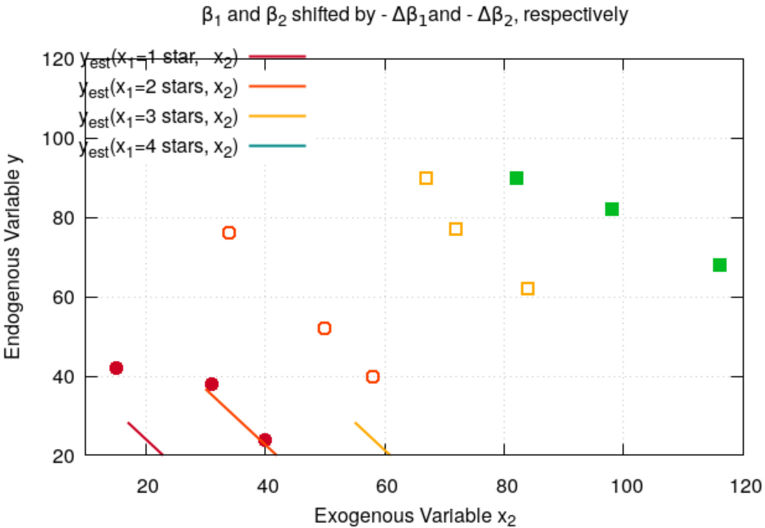
Effect of mis-fit parameters II: small effect if β_1 and β_2 have opposite misfits



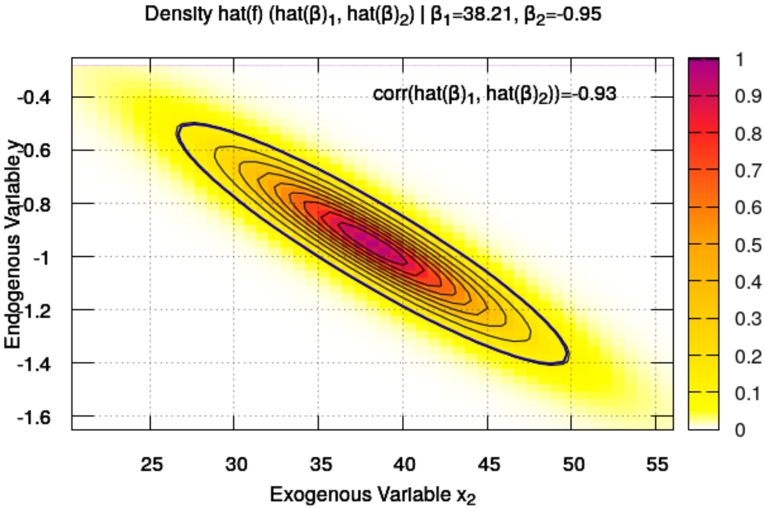
Effect of mis-fit parameters III: large effect if β_1 and β_2 have both positive misfits



Effect of mis-fit parameters IV: large effect if β_1 and β_2 have both negative misfits



All this results in a negative correlation between the estimation errors for β_1 and β_2



Special case 1: No exogenous variables

- Model: $y = \beta_0 + \epsilon := \mu + \epsilon$
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$$(\mathbf{X}'\mathbf{X})^{-1} = \frac{1}{n}, \quad \mathbf{X}'\boldsymbol{y} = \sum_i y_i = n\bar{y},$$

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- Model (with $x_1 = x$): $y = \beta_0 + \beta_1 x + \epsilon$
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$$\mathbf{X} = \begin{pmatrix} 1 & x_1 \\ \vdots & \vdots \\ 1 & x_n \end{pmatrix}, \quad \mathbf{X}'\mathbf{X} = \begin{pmatrix} n & n\bar{x} \\ n\bar{x} & \sum x_i^2 \end{pmatrix}$$

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- Variance-covariance matrix (assuming w/o loss of generality $\bar{x} = 0$):

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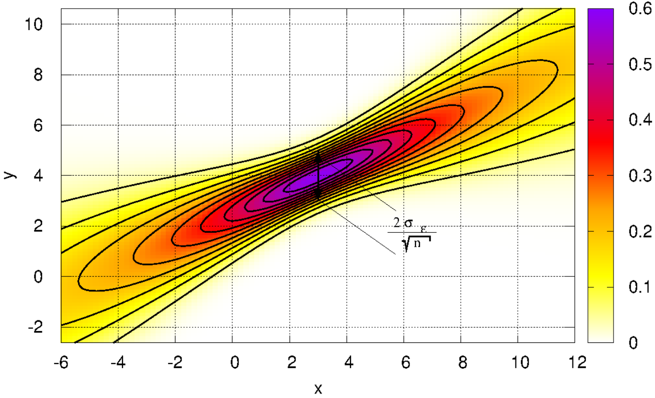
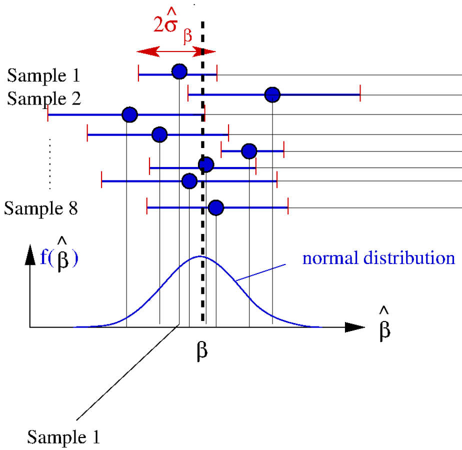
Probability density for $\hat{y}(x)$ for simple linear regression



- If the Gauß-Markov assumptions apply, the model estimation errors $\hat{y}(x) - y(x)$ are Gaussian distributed
- The expectation and variance depends on x; the standard error is hyperbola-shaped.

3.2. Confidence Intervals: where the Student-t distribution comes from



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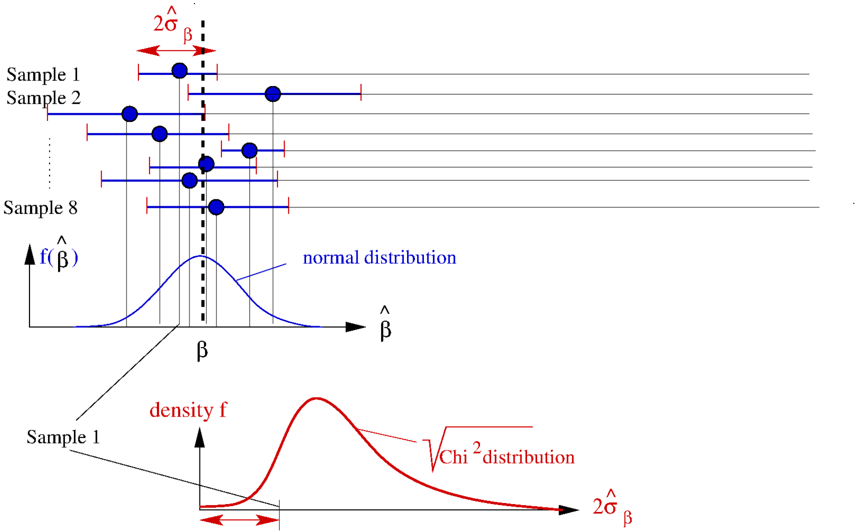

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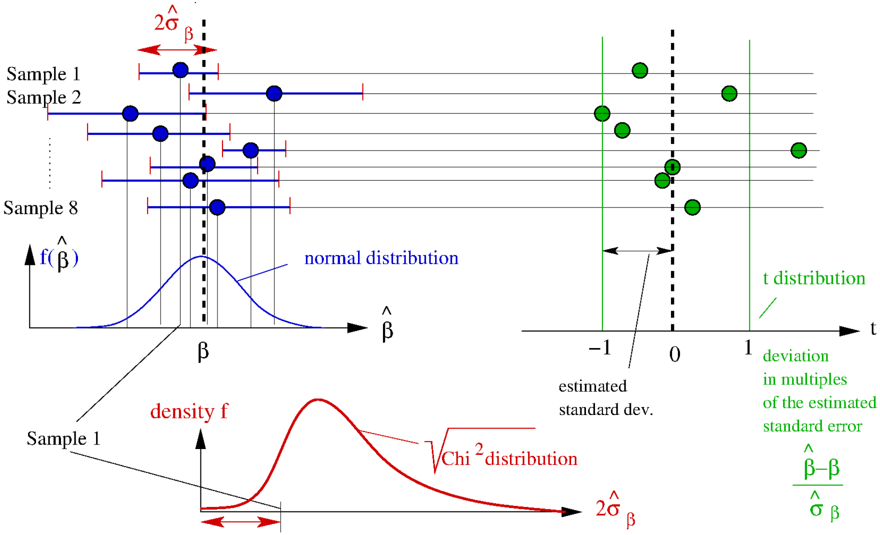

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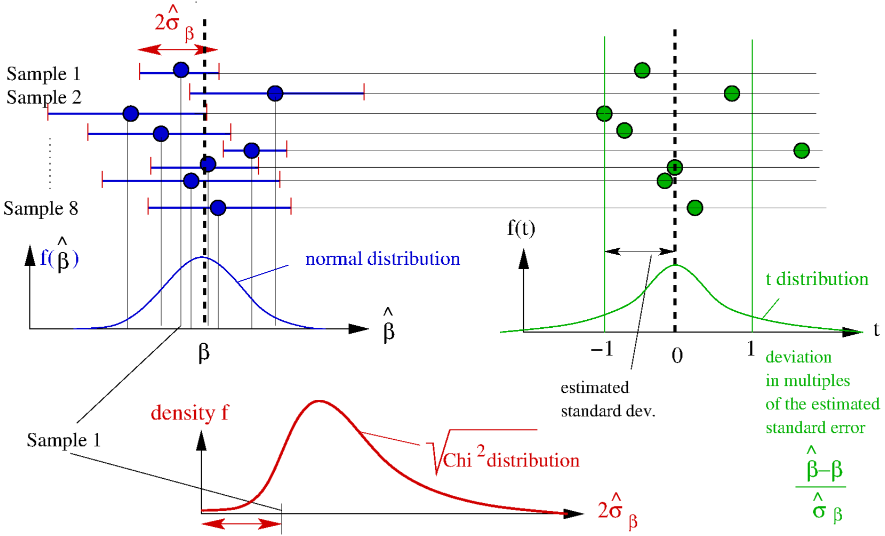
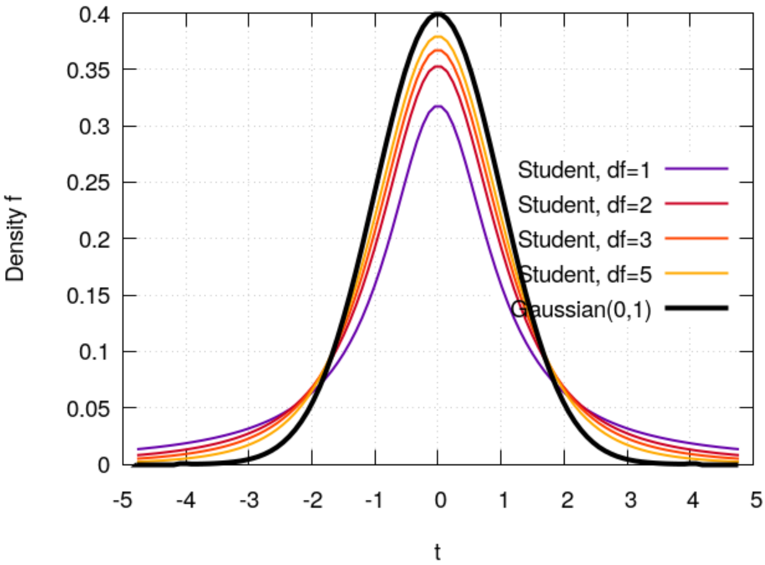


Densities of standard normal vs. Student-t distribution



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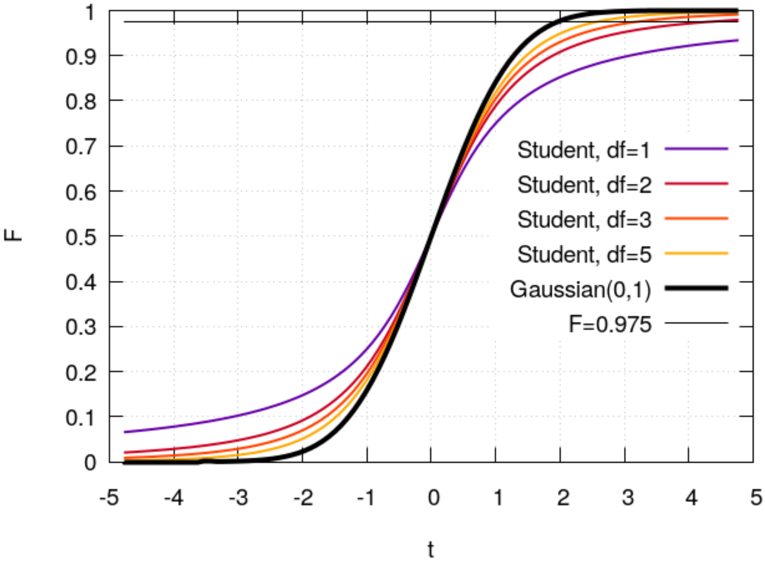

Calculation of the confidence intervals (CI)

$$\mathrm{CI}_{\beta_j}^{(\alpha)} : \beta_j \in \left[\hat{\beta}_j - \Delta\hat{\beta}_j, \hat{\beta}_j + \Delta\hat{\beta}_j \right], \quad \Delta\hat{\beta}_j = t_{1-\alpha/2}^{(n-J-1)} \hat{\sigma}_{\hat{\beta}_j}.$$

- $t_{1-\alpha/2}$: Quantile (inverse of) the distribution function
- CI "uncertainty principle": Higher sensitivity implies higher α error.

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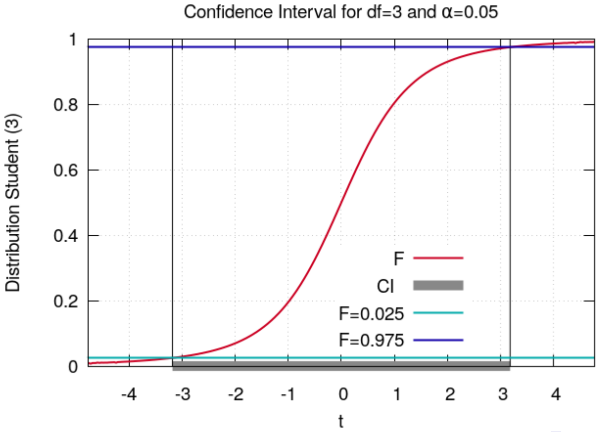

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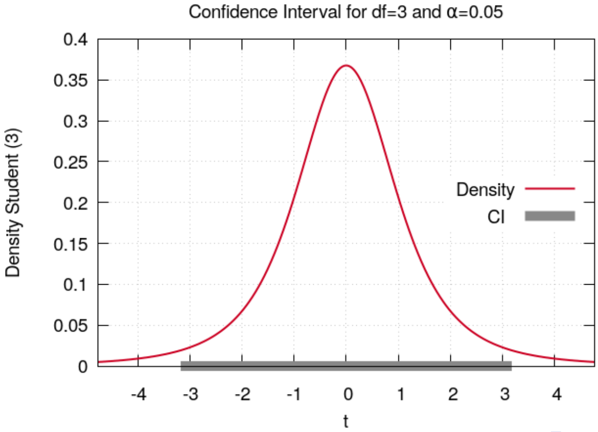
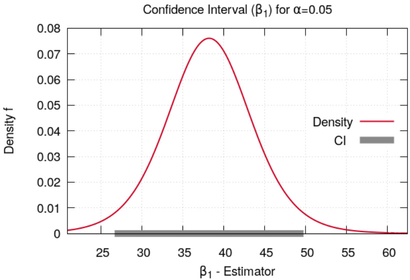


Hotel example: CI for the appraisal for "stars" β_1 (full model)



Model: $y(\boldsymbol{x}) = \sum_j \beta_j x_j + \epsilon$

Factors:
$x_0 = 1$, x_1: #stars, x_2: price

Confidence interval (CI):

$$\beta_1 \in \left[\hat{\beta}_1 - \Delta\hat{\beta}_1^{(\alpha)}, \hat{\beta}_1 + \Delta\hat{\beta}_1^{(\alpha)}\right]$$

$$\Delta\hat{\beta}_1^{(\alpha)} = t_{1-\alpha/2}^{(n-3)} \sqrt{\hat{V}(\hat{\beta}_1)}$$

$$\hat{V}(\hat{\beta}_1) = \hat{\sigma}_\epsilon^2 \left[(\mathbf{X}'\mathbf{X})^{-1}\right]_{11}$$

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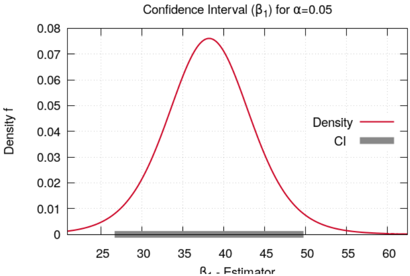

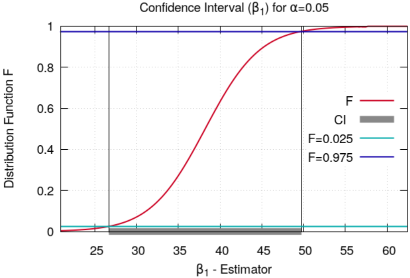

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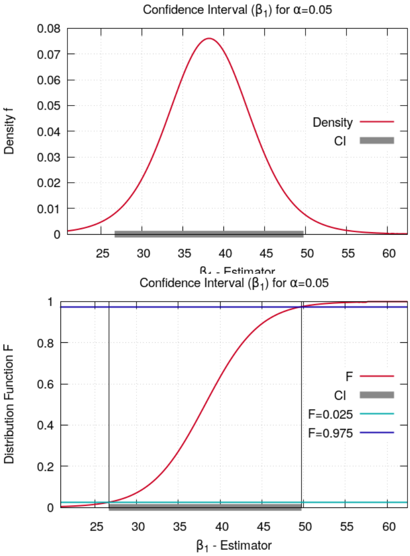

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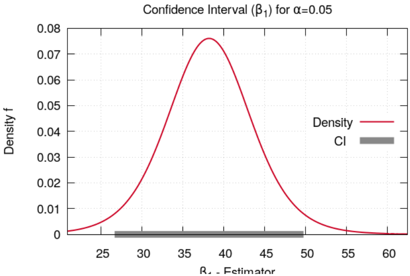

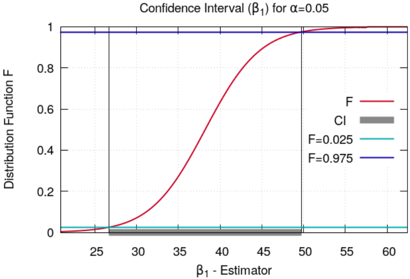

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